SPECTRAL DECOMPOSITION OF DISPERSION MATRIX FOR THE MIXED ANALYSIS OF VARIANCE MODEL

B. GÜVEN

Department of Statistics, Middle East Technical University 06531, Ankara Turkey.

(Received June 27, 2000; Accepted Oct. 11, 2000)

ABSTRACT

The spectral decomposition of the variance-covariance matrix for a balanced mixed analysis of variance model is presented. The model consists of crossed and/or nested factors with either replicated or nonreplicated.

1. INTRODUCTION

The spectral decomposition of a variance-covariance matrix (dispersion matrix) V is useful for finding its powers V^{α} where α is any real number. In particular, $\alpha = -1$, V^{-1} is useful for estimation or $\alpha = -\frac{1}{2}$, $V^{-1/2}$ is useful for the transforming a linear model to a model with i.i.d. error terms.

The problem has been discussed before by Searle and Henderson [1] and Wansbeek and Kapteyn [2]. In both studies, it is supposed that the form of the spectral decomposition of V is of the same form of V. Then they obtained idempotent matrices in the spectral decomposition of V by equating V and its assumed spectral decomposition.

However our solution is based on deriving an idempotent matrix from eigenvectors for the corresponding eigenvalue in the spectral decomposition of ${\bf V}$ without assuming any form of the spectral decomposition of ${\bf V}$.

2. THE DISPERSION MATRIX

The variance-covariance matrix V for a balanced k-factor mixed analysis of variance model is of the following structure

$$V = \sum_{d} \lambda_{d} N_{d}$$
 (1)

118 B. GÜVEN

where **d** is a k-vector of zeros and ones. The summation is taken over 2^k -elements. The λ_d are nonnegative parameters. Let $\mathbf{d}=(i_1,i_2,...,i_k)$ with i_r or 1 for r=1,2...k. Then the matrix $\mathbf{N_d}$ in (1) are defined as

$$\mathbf{N_d} = \mathbf{J_1^{i_1}} \otimes \mathbf{J_2^{i_2}} \otimes ... \otimes \mathbf{J_k^{i_k}}$$

with $\mathbf{J}_{\tau}^{0} = \mathbf{I}_{\tau}$, where \mathbf{J}_{τ} and \mathbf{I}_{τ} are respectively a matrix of ones and an identity matrix of order \mathbf{n}_{τ} for r=1,2,...,k, the symbol \otimes denotes the Kronecker product of matrices.

The full rank of (1), leading that all eigenvalues of (1) are nonzero, is provided by $N_{00\dots00} = I_{12\dots k}$ where $I_{12\dots k}$ is an identity matrix of order $\prod_{r=1}^k n_r$. A linear space generated by the columns of (1) is the sum of linear subspaces generated by the columns of 2^k matrices of order $\prod_{r=1}^k n_r$ given by $I_{12\dots k}, N_{00\dots01}, \dots, N_{11\dots11}$ and then is spanned by the set of basis $I_{12\dots k}, N_{00\dots01}, \dots, N_{11\dots11}$. However the basis for any $N_{i_1i_2\dots i_k}$ are the linear combination of the basis for $I_{12\dots k}$. As a result, a linear space generated by the columns of (1) is spanned by the set of basis for $I_{12\dots k}$.

3. THE SPECTRAL DECOMPOSITION

Let

Let $t_1, t_2, ..., t_k$ be denoted by t with $t_r = 0$ or 1. The 2^k (possibly) distinct eigenvalues of (1) given by [1] are:

$$\phi_{t} = \sum_{d} \lambda_{d} x_{t_{1}}^{i_{1}} x_{t_{2}}^{i_{2}} ... x_{t_{k}}^{i_{k}}$$
 (2)

with multiplicity $\prod_{r=1}^k (n_r-1)^{1-t_r}$ and where $\mathbf{x}_{t_r}^{i_r}$ is the eigenvalue of the matrix $\mathbf{J}_r^{i_r}$ given by

$$\mathbf{x}_{\mathbf{t_r}}^{\mathbf{i_r}} = \begin{cases} 0 & \text{if} \quad \mathbf{t_r} = 0\\ \mathbf{n_r} & \text{if} \quad \mathbf{t_r} = 1 \end{cases}$$

with multiplicity $(n_r - 1)^{1-t_r}$ if $i_r = 1$. $x_{t_r}^{i_r} = 1$ for $t_r = 0,1$ with multiplicity n_r if $i_r = 0$. An eigenvector v_{t_r} for $x_{t_r}^{i_r}$ will be:

$$\mathbf{v}_{t_{r}} = \begin{cases} \mathbf{\xi}_{\mathbf{n_{r}}k} & k = 1, 2, ..., \mathbf{n_{r}} - 1 & \text{if} \quad t_{r} = 0\\ \frac{1}{\sqrt{\mathbf{n_{r}}}} \mathbf{1}_{r} & \text{if} \quad t_{r} = 1 \end{cases}$$
(3)

and $\xi_{n_r 1}, \xi_{n_r 2}, ..., \xi_{n_r n_r - 1}, \frac{1}{\sqrt{n_r}} \mathbf{1}_r$ is an orthonormal set, $\mathbf{1}_r$ is a $n_r \times 1$ vector of ones.

$$\mathbf{v_t} = \mathbf{v_{t_1}} \otimes \mathbf{v_{t_2}} \otimes ... \otimes \mathbf{v_{t_k}}. \tag{4}$$

Then v_t is an eigenvector for ϕ_d in (2) since $J_r^{i_T} v_{t_r} = x_{t_r}^{i_T} v_{t_r}$ and

$$\begin{split} \mathbf{N_{d}v_{t}} &= (\mathbf{J}_{1}^{i_{1}} \otimes \mathbf{J}_{2}^{i_{2}} \otimes ... \otimes \mathbf{J}_{k}^{i_{k}}) (\mathbf{v_{t_{1}}} \otimes \mathbf{v_{t_{2}}} \otimes ... \otimes \mathbf{v_{t_{k}}}) \\ &= \mathbf{J}_{1}^{i_{1}} \mathbf{v_{t_{1}}} \otimes \mathbf{J}_{2}^{i_{2}} \mathbf{v_{t_{2}}} \otimes ... \otimes \mathbf{J}_{k}^{i_{k}} \mathbf{v_{t_{k}}} \\ &= \mathbf{x}_{t_{1}}^{i_{1}} \mathbf{v_{t_{1}}} \otimes \mathbf{x}_{t_{2}}^{i_{2}} \mathbf{v_{t_{2}}} \otimes ... \otimes \mathbf{x}_{t_{k}}^{i_{k}} \mathbf{v_{t_{k}}} = (\mathbf{x}_{t_{1}}^{i_{1}} \mathbf{x}_{t_{2}}^{i_{2}} ... \mathbf{x}_{t_{k}}^{i_{k}}) (\mathbf{v_{t_{1}}} \otimes \mathbf{v_{t_{2}}} \otimes ... \otimes \mathbf{v_{t_{k}}}) \,. \end{split}$$

Consequently,

$$V_{\nu_t} = \sum_{d} \lambda_d N_d \nu_t = \phi_t \nu_t \ .$$

Let $\mathbf{P_{t_r}} = \mathbf{v_{t_r}} \mathbf{v_{t_r}'}$ for r=1,2,...,k where $\mathbf{v_{t_r}}$ in (3) is an eigenvector of $\mathbf{J_r^{i_r}}$ and $\mathbf{M_t} = \mathbf{v_t} \mathbf{v_t'}$ where $\mathbf{v_t}$ in (4) is an eigenvector of (1). Then both $\mathbf{P_{t_r}}$ and $\mathbf{M_t}$ are idempotent matrices and the spectral decomposition of (1) is:

$$V = \sum_t \phi_t M_t$$

where

$$\mathbf{M}_{\mathbf{t}} = \mathbf{P}_{\mathbf{t}_1} \otimes \mathbf{P}_{\mathbf{t}_2} \otimes \dots \otimes \mathbf{P}_{\mathbf{t}_r} \tag{5}$$

with

$$\mathcal{F}_{\mathbf{t_r}} = \begin{cases} \frac{n_r - 1}{\sum_{\ell=1}^{n_r \ell} \xi_{\mathbf{n_r}\ell}'} & \text{if} \quad \mathbf{t_r} = 0\\ \frac{1}{n_r} J_r & \text{if} \quad \mathbf{t_r} = 1 \end{cases}$$

$$(6)$$

Consider a matrix $\mathbf{I_r} + \mathbf{J_r}$ having eigenvalues 1 with multiplicity $\mathbf{n_r} - 1$ and $1 + \mathbf{n_r}$ and the respective orthonormal eigenvectors $\xi_{\mathbf{n_r}\ell}$, $\ell = 1,2,...,\mathbf{n_r} - 1$ for 1 and $1/\sqrt{\mathbf{n_r}} \cdot \mathbf{I_r}$. Then the spectral decomposition of $\mathbf{I_r} + \mathbf{J_r}$ is:

$$I_{r} + J_{r} = \sum_{\ell=1}^{n_{r}-1} \xi_{n_{\ell}} \xi_{n_{r}\ell}^{\ell} + (1 + n_{r}) \frac{1}{n_{r}} J_{r}$$
 (7)

Using (7),(6) can be rewritten as:

$$\mathbf{P}_{\mathbf{t}_{r}} = \begin{cases} \mathbf{I}_{r} - \frac{1}{n_{r}} \mathbf{J}_{r} & \text{if} \quad \mathbf{t}_{r} = 0\\ \\ \frac{1}{n_{r}} \mathbf{J}_{r} & \text{if} \quad \mathbf{t}_{r} = 1 \end{cases}$$
(8)

where the rank of P_{t_r} is $(n_r - 1)^{1-t_r}$. From (5) with (8), it can be seen that M_t has rank $\sum_{k=1}^{k} (n_k - 1)^{1-t_r}$ and $M_t M_{t^*} = 0$ for $t \neq t^*$.

Consider a mixed model representing an experiment that is replicated n_k -times. (1) can be rewritten as

$$\mathbf{V} = \lambda_{00...00} \mathbf{I}_{\mathbf{n}} + \sum_{\mathbf{d}} \lambda_{\mathbf{d}} \mathbf{N}_{\mathbf{d}} . \tag{9}$$

since $\lambda_{00...00}$ is positive and $\lambda_{t_1t_2...t_{k-1}0}$ is zero for at least one of nonzero t_r where r=1,2,...,k-1. Here $\mathbf{d}=(i_1i_2...i_{k-1})$ with $i_r=0$ for r=0,1,...k-1, the summation on the right hand side of (9) is taken over 2^{k-1} -elements and

$$N_d = J_1^{i_1} \otimes J_2^{i_2} \otimes ... \otimes J_{k-1}^{i_{k-1}} \otimes J_k.$$

It follows that (2) can be:

$$\phi_{t} = \lambda_{00...00} \sum_{d} \lambda_{d} x_{1}^{i_{1}} x_{2}^{i_{2}} ... x_{k-1}^{i_{k-1}} x_{k}$$
(10)

From (10), $\phi_{t_1t_2...t_{k-1}0} = \lambda_{00...00}$ with multiplicity

$$(n_k-l)\sum_{t_1t_2\dots t_{k-1}=0}^{l} \ \prod_{r=1}^{k-l} (n_r-l)^{l-t_r} = (n_k-l)\prod_{r=1}^{k-l} n_r.$$

So 2^{k-1} -eigenvalues of (9) are the same and equal to the smallest eigenvalue $\lambda_{00\dots 00}$. The corresponding idempotent matrix for $\lambda_{00\dots 00}$ will be

$$\mathbf{M}_{k-1} \otimes (\mathbf{I}_r - \frac{1}{n_r} \mathbf{J}_r)$$

where a matrix \mathbf{M}_{k-1} of order $\prod_{r=1}^{k-1} \mathbf{n}_r$ is:

$$\mathbf{M}_{k-1} = \sum_{\substack{t_1t_2...t_{k-1}=0}}^{1} \mathbf{P}_{t_1} \otimes \mathbf{P}_{t_2} \otimes ... \otimes \mathbf{P}_{t_k-1}$$

with

$$\begin{split} \text{rank}(\boldsymbol{M}_{k-1}) &= \sum_{t_1 t_2 \dots t_{k-1} = 0}^{1} \text{rank}(\boldsymbol{P}_{t_1}) \text{rank}(\boldsymbol{P}_{t_2}) \dots \text{rank}(\boldsymbol{P}_{t_{k}-1}) \\ &= \sum_{t_1 t_2 \dots t_{k-1} = 0}^{1} \prod_{r=1}^{k-1} (n_r - 1)^{l - t_r} = \prod_{r=1}^{k-1} \sum_{t_r = 0}^{1} (n_k - 1)^{l - t_r} = \prod_{r=1}^{k-1} n_r \,. \end{split}$$

Then M_{k-1} is an identity matrix since the full rank idempotent matrix is unique and equal to an identity matrix.

The eigenvalue of (9) is of the form

$$\phi_{t_1t_2...t_{k-1}}^* = \phi_{t_1t_2...t_{k-1}} = \lambda_{00...00} + n_k \sum_{i} \lambda_{d} x_{t_1}^{i_1} x_{t_2}^{i_2} ... x_{t_{k-1}}^{i_{k+1}}.$$
 (11)

Then the spectral decomposition of (9) according to $2^{k-1}+1$ (possibly) distinct eigenvalues of (9) is:

$$V = \lambda_{00...00} \mathbf{I}_{12...k-1} \otimes (\mathbf{I}_{k} - \frac{1}{n_{k}} \mathbf{J}_{k})$$

$$+ \sum_{t_{1}t_{2}...t_{k}-1}^{1} \phi_{t_{1}t_{2}...t_{k}-1}^{*} \mathbf{P}_{t_{1}} \otimes \mathbf{P}_{t_{2}} \otimes ... \otimes \mathbf{P}_{t_{k}-1} \otimes \frac{1}{n_{k}} \mathbf{J}_{k}$$
(12)

where $I_{12...k-1}$ is an identity matrix of order $\prod_{r=1}^{k-1} n_r$ and P_{t_r} is in (8).

4. CONCLUDING REMARKS

The spectral decomposition of V provides easily the computation of V^{α} for any real α since

$$V^\alpha = \sum_t \phi_t^\alpha M_t$$

where ϕ_t and M_t are defined in (2) and (5) respectively.

The nonnegative parameters λ_d in (1) correspond to variance components. From (2), the eigenvalues of the variance-covariance matrix is the linear combination of variance components. It is not necessary to recompute the spectral decomposition of V^* where a new variance-covariance matrix V^* is obtained by removing some λ_d 's from V. It can be obtained by removing the corresponding λ_d 's from the spectral decomposition of V.

A half number of eigenvalues of V are same and equal to the smallest eigenvalue when there is a replication. In this case, the summations in both (11) and (12) are taken over 2^{k-1} , instead of 2^k . This facilitates the computation of the spectral decomposition of V.

5. AN EXAMPLE

Consider the two-way random effect model $y_{ijk} = \mu + a_i + b_j + c_{ij} + e_{ijk}$, $i = 1,2,...,n_1$, $j = 1,2,...,n_2$, $k = 1,2,...,n_3$ where

$$\boldsymbol{a}_i \sim N(0, \sigma_a^2), \, \boldsymbol{b}_j \sim N(0, \sigma_b^2), \, \boldsymbol{c}_{ij} \sim N(0, \sigma_c^2), \, \boldsymbol{e}_{ijk} \sim N(0, \sigma_e^2)$$

and they are independent. The variance-covariance matrix for this model is:

$$V = \sigma_a^2 \mathbf{I}_1 \otimes \mathbf{J}_2 \otimes \mathbf{J}_3 \otimes \sigma_b^2 \mathbf{J}_1 \otimes \mathbf{I}_2 \otimes \mathbf{J}_3 \otimes \sigma_c^2 \mathbf{I}_1 \otimes \mathbf{I}_2 \otimes \mathbf{I}_3 \otimes \sigma_e^2 \mathbf{I}_1 \otimes \mathbf{I}_2 \otimes \mathbf{I}_3$$

where I_r and J_r for r=1,2,3 are a $n_r \times n_r$ identity matrix and a $n_r \times n_r$ matrix of ones respectively.

Define $\lambda_{000} = \sigma_e^2$, $\lambda_{001} = \sigma_c^2$, $\lambda_{011} = \sigma_a^2$, $\lambda_{101} = \sigma_b^2$ and the other λ_d 's are zero. The 2^3 eigenvalues of V are:

$$\begin{split} \varphi_{t_1 t_2 t_3} &= \sum_{i_1 i_2 i_3 = 0}^{1} \lambda_{i_1 i_2 i_3} x_{t_1}^{i_1} x_{t_2}^{i_2} x_{t_3}^{i_3} \\ &= \lambda_{000} + \lambda_{001} x_{t_3} + \lambda_{011} x_{t_2} x_{t_3} + \lambda_{101} x_{t_1} x_{t_3} \end{split}$$

where $x_{t_r} = 0$ if $t_r = 0$, $x_{t_r} = n_r$ if $t_r = 1$ for r=1,2,3. Then

$$\phi_{000} = \phi_{010} = \phi_{100} = \phi_{110} = \lambda_{000} \; , \; \phi_{001} = \lambda_{000} + n_3 \lambda_{001} , \\ \phi_{011} = \lambda_{000} + n_3 \lambda_{001} + n_2 n_3 \lambda_{011} \; , \\ \phi_{010} = \lambda_{010} + n_3 \lambda_{011} + n_2 n_3 \lambda_{011} \; , \\ \phi_{010} = \lambda_{010} + n_3 \lambda_{011} + n_2 n_3 \lambda_{011} \; , \\ \phi_{010} = \lambda_{010} + n_3 \lambda_{011} + n_2 n_3 \lambda_{011} \; , \\ \phi_{010} = \lambda_{010} + n_3 \lambda_{011} + n_2 n_3 \lambda_{011} \; , \\ \phi_{010} = \lambda_{010} + n_3 \lambda_{011} + n_2 n_3 \lambda_{011} \; , \\ \phi_{010} = \lambda_{010} + n_3 \lambda_{011} + n_3 \lambda_{011} \; , \\ \phi_{011} = \lambda_{010} + n_3 \lambda_{011} + n_3 \lambda_{011} \; , \\ \phi_{011} = \lambda_{010} + n_3 \lambda_{011} + n_3 \lambda_{011} \; , \\ \phi_{011} = \lambda_{010} + n_3 \lambda_{011} + n_3 \lambda_{011} + n_3 \lambda_{011} \; , \\ \phi_{011} = \lambda_{010} + n_3 \lambda_{011} + n_3 \lambda_{011} + n_3 \lambda_{011} \; , \\ \phi_{011} = \lambda_{010} + n_3 \lambda_{011} + n_3 \lambda_{011} + n_3 \lambda_{011} \; , \\ \phi_{011} = \lambda_{010} + n_3 \lambda_{011} \; , \\ \phi_{011} = \lambda_{010} + n_3 \lambda_{011} + n_$$

$$\phi_{101} = \lambda_{000} + n_3 \lambda_{001} + n_1 n_3 \lambda_{101}, \ \phi_{111} = \lambda_{000} + n_3 \lambda_{001} + n_1 n_3 \lambda_{101} + n_2 n_3 \lambda_{011}.$$

The spectral decomposition of V is:

122 B. GÜVEN

$$\begin{split} \mathbf{V} &= \phi_{000} \, \mathbf{I}_{12} \otimes (\mathbf{I}_3 - \frac{1}{n_3} \mathbf{J}_3) + \phi_{001} \, \mathbf{I}_1 \otimes \mathbf{I}_2 \otimes \frac{1}{n_3} \mathbf{J}_3 \\ &+ \phi_{011} (\mathbf{I}_1 - \frac{1}{n_1} \mathbf{J}_1) \otimes \frac{1}{n_2} \mathbf{J}_2 \otimes \frac{1}{n_3} \mathbf{J}_3 \\ &+ \phi_{101} \frac{1}{n_1} \mathbf{J}_1 \otimes (\mathbf{I}_2 - \frac{1}{n_2} \mathbf{J}_2) \otimes \frac{1}{n_3} \mathbf{J}_3 \\ &+ \phi_{111} \frac{1}{n_1} \mathbf{J}_1 \otimes \frac{1}{n_2} \mathbf{J}_2 \otimes \frac{1}{n_3} \mathbf{J}_3. \end{split}$$

Acknowledgment

The author thanks the referee for making comments which clearify the paper.

REFERENCES

- [1] R. SEARLE, and H. V. HENDERSON. Dispersion Matrices for Variance Components Models, Journal of American Statistical Association, 74 (1979), 465-470.
- [2] T. WANSBEEK and A. KAPTEYN, A Note on Spectral Decomposition and Maximum Likelihood Estimation in ANOVA Models with Balanced Data, Statistics & Probability Letters, 1 (1983), 213-215.