

COMMUNICATIONS

DE LA FACULTÉ DES SCIENCES
DE L'UNIVERSITÉ D'ANKARA

Série A₁: Mathématiques

TOME : 33

ANNÉE : 1984

On Truncated Poisson Queue With Balking, Reneging and Heterogeneity.

by

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Ankara, Turquie

Communications de la Faculté des Sciences de l'Université d'Ankara

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TURQUIE

On Truncated Poisson Queue With Balking, Reneging and Heterogeneity.

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(Received February 21, 1984 and accepted August 13, 1984)

ABSTRACT

In this paper the author treats a poissonian queue with the following three properties: Balking, Reneging and Heterogeneity. A modified queue discipline to the classical one $1st$ in $1st$ out (FIFO) is used with a more general condition. In fact this paper is a continuation and more generalization to the works done by many researchers such as: Abou - El - Ata, Singh, Krishna - moorthi and others.

1. INTRODUCTION

In fact this work was early studied by many researchers in case of only one property. Krishnamoorthi [3] considered a poisson queue with heterogeneity and two alternative queue disciplines one with slight and the other with a great modification of the classical one (FIFO). Singh [5] analyzed a Markovian queue with balking and heterogeneity with a modified queue discipline but different than those of Krishnamoorthi. Abou - El - Ata [1] considered a Markovian queue with both balking and heterogeneity with a modified queue discipline of both Singh [5] and Krishnamoorthi [3]. In fact he had generalized both of their works.

In this paper the author adds a more third property $(n-1)\alpha$ due to the reneging factor. This means that the author treats a truncated poisson queue (capacity N) with balking, reneging and heterogeneity. A modified queue discipline of both Singh and Krishnamoorthi is used with a more general condition. In fact and really this work is a generalization of most of the previous works.

2. Analyzing of the problem

Consider the truncated two-channels queue: $M/M/2/N$ with balking, reneging and heterogeneity. The arrivals are poisson distributed with mean

rate λ and the service times are exponentially distributed with mean rates $\mu_i, i = 1,2$. The queue discipline considered is:

- (i) If both channels are free, the top unit of the queue goes to channel I with probability π_1 or to channel II with probability π_2 ($\pi_1 + \pi_2 = 1$).
- (ii) If one of them is free, the top unit goes to it.
- (iii) If both channels are busy the top unit waits until the 1st one is free and then goes to it. The new coming unit joins the queue with probability β (i.e the unit balks with prob. $1-\beta$). Any unit in the queue leaves it with probability $[(n-2)\alpha], n = 2 (1)N$, {which means it reneges with prob. $(n-2) \alpha$ }.

Let us define the probabilities:

- $p_0(t) = p_{00}(t) = \text{prob. \{there is no unit in the system at time t\}}$
- $p_{10}(t) = \text{prob. \{there is one unit in channel I at time t\}}$
- $p_{01}(t) = \text{prob. \{there is one unit in channel II at time t\}}$
- $p_n(t) = \text{prob. \{there are n units in the system at time t, n = 2 (1) \dots\}}$

Also: $p_1(t) = p_{10}(t) + p_{01}(t)$, and; $p_2(t) = p_{11}(t)$

As in the usual arguments of the δ -technique the steady-state probability difference equations are:

$$-\lambda p_0 + \mu_1 p_{10} + \mu_2 p_{01} = 0 \quad n = 0 \quad (1)$$

$$\left. \begin{aligned} -(\lambda + \mu_1) p_{10} + \mu_2 p_{11} + \lambda \pi_1 p_0 &= 0 \\ -(\lambda + \mu_2) p_{01} + \mu_1 p_{11} + \lambda \pi_2 p_0 &= 0 \end{aligned} \right\} n = 1 \quad (2)$$

$$-(\beta\lambda + \mu) p_2 + (\mu + \alpha) p_3 + \lambda p_1 = 0 \quad n = 2 \quad (3)$$

$$-[\beta\lambda + \mu + (n-2)\alpha]P_n + [\mu + (n-1)\alpha]P_{n+1} + \beta\lambda P_{n-1} = 0, \quad N-1 \geq n > 2 \quad (4)$$

$$-[\mu + (N-2)\alpha] p_N + \beta\lambda p_{N-1} = 0 \quad n = N \quad (5)$$

Where: $\mu = \mu_1 + \mu_2$.

Adding the 1st four equations in (1), (2) and (3)

We get:

$$(\mu + \alpha) p_3 - \beta\lambda p_2 = 0 \quad (6)$$

From (3) and (6) we have:

$$0 = (\mu + \alpha) p_3 - \beta\lambda p_2 = \mu p_2 - \lambda p_1$$

$$\text{i.e.} \quad p_2 = \rho p_1 \quad (7)$$

From: (2) and (7) we find:

$$p_{10} = \frac{\lambda [\lambda + (\mu_1 + \mu_2) \pi_1]}{\mu_1 [2\lambda + \mu_1 + \mu_2]} p_0, \quad \text{and};$$

$$p_{01} = \frac{\lambda [\lambda + (\mu_1 + \mu_2) \pi_2]}{\mu_2 [2\lambda + \mu_1 + \mu_2]} p_0$$

$$\text{Therefore: } p_1 = \Delta p_0 \quad (8)$$

$$\begin{aligned} \text{Where: } \Delta &= \frac{\lambda (\mu_1 + \mu_2) [\lambda + \mu_1 \pi_2 + \mu_2 \pi_1]}{\mu_1 \mu_2 (2\lambda + \mu_1 + \mu_2)} \\ &= \frac{\lambda [\lambda + \mu_1 \pi_2 + \mu_2 \pi_1]}{\mu_1 \mu_2 (2\rho + 1)}, \quad \rho = \frac{\lambda}{\mu} \end{aligned} \quad (9)$$

From (4) and (6) we can get:

$$\begin{aligned} [\mu + (n-1)\alpha] p_{n+1} - \beta\lambda p_n &= [\mu + (n-2)\alpha] p^n - \beta\lambda p_{n-1} \\ &= \dots = (\mu + \alpha) p_3 - \beta\lambda p_2 = 0 \end{aligned}$$

$$\text{i.e. } p_n = \frac{\beta\lambda}{\mu + (n-2)\alpha} p_{n-1}, \quad n = 2 \text{ (1) } N-1 \quad (10)$$

Using (7) in (10) we find:

$$p_n = \frac{\rho (\beta\gamma)^{n-2}}{(\delta + 1)_{n-2}} p_1, \quad n = 2 \text{ (1) } N-1 \quad (11)$$

$$\text{Where: } \gamma = \frac{\lambda}{\alpha} \quad \text{and} \quad \delta = \frac{\mu}{\alpha}$$

$$\text{Also from (5): } p_N = \frac{\beta\gamma}{\delta + (N-2)} p_{N-1} \quad (12)$$

Thus from (8), (11), and (12) we have:

$$p_n = \frac{\Delta\rho (\beta\gamma)^{n-2}}{(\delta+1)_{n-2}} p_0 \quad n = 2 \text{ (1) } N \quad (13)$$

where p_0 could be found from the boundary condition:
as follows:

$$\sum_{n=0}^N p_n = 1$$

$$p_0 + p_1 + \sum_{n=2}^N p_n = 1$$

$$p_0 + \Delta p_0 + \Delta \rho p_0 \sum_{n=0}^N \frac{(\beta\gamma)^{n-2}}{(\delta+1)_{n-2}} = 1$$

$$\text{i.e } p_0 + \Delta p_0 + \Delta \rho p_0 \sum_{n=0}^{N-2} \frac{(1)_n (\beta\gamma)^n}{n! (\delta+1)_n} = 1$$

$$\text{Thus: } p_0^{-1} = 1 + \Delta + \Delta \rho {}_1F_1 (-(N-2); \delta+1; -\beta\gamma) \quad (14)$$

where Δ is given in (9).

The measures of effectiveness are:

$$\begin{aligned} E(n) &= \sum_{n=1}^N n p_n = \Delta p_0 + \Delta \rho p_0 \sum_{n=2}^N \frac{n (\beta\gamma)^{n-2}}{(\delta+1)_{n-2}} \\ &= \Delta p_0 + \Delta \rho \beta \gamma p_0 \frac{d}{d(\beta\gamma)} \{ {}_1F_1 (-(N-2); \delta+1, -\beta\gamma) \} \\ &\quad + 2 \Delta \rho p_0 {}_1F_1 (-(N-2); \delta+1; -\beta\gamma) \end{aligned} \quad (15)$$

$$\begin{aligned} \text{and } E(m) &= \sum_{n=2}^N (n-1) p_n \\ &= \Delta \rho \beta \gamma p_0 \frac{d}{d(\beta\gamma)} \{ {}_1F_1 (-(N-2); \delta+1; \beta\gamma) \} \\ &\quad + \Delta \rho p_0 {}_1F_1 (-(N-2); \delta+1, -\beta\gamma) \end{aligned} \quad (16)$$

Corollary (1): Let $\alpha = 0$, this is the truncated poisson queue with both balking and heterogeneity. Thus from (13) we get

$$p_n = \Delta \rho (\beta \rho)^{n-2} \cdot p_0, \quad n = 2 (1) N \quad (17)$$

$$\text{Where: } p_0^{-1} = 1 + \frac{\Delta}{1 - \beta\rho} [1 + (1 - \beta) \rho - \rho (\beta\rho)^{N-1}] \quad (18)$$

and Δ is given in (9).

The measures of effectiveness in this case are:

$$E(n) = \Delta p_0 (1 + x) \text{ and } E(m) = \Delta p_0 x \tag{19}$$

$$\text{where } x = \frac{\beta \{1 - (\beta\rho)^{N-1}\}}{1 - \beta\rho} + \frac{\rho \{1 - (N-1)(\beta\rho)^{N-2} + N(\beta\rho)^{N-1}\}}{(1 - \beta\rho)^2} \tag{20}$$

Corollary (2): Let $\alpha = 0$ and $N \rightarrow \infty$, we have the poisson queue with both balking and heterogeneity which is Abou - El - Ata [1].

Thus from (7), (18), (19) and (20) we find:

$$p_n = \Delta \rho (\beta\rho)^{n-2} \cdot p_0, \tag{21}$$

$$p_0 = 1 + \frac{\Delta [1 + (1 - \beta) \rho]}{1 - \beta\rho} \\ = \frac{\lambda [\lambda + \mu_1 \pi_2 + \mu_2 \pi_1] [1 + (1 - \beta) \rho]}{\mu_1 \mu_2 (2 \rho + 1) (1 - \beta\rho)} \tag{22}$$

$$E(n) = \Delta p_0 \cdot \left[1 + \frac{\rho}{1 - \beta\rho} + \frac{\rho}{(1 - \beta\rho)^2} \right] \tag{23}$$

$$\text{and } E(m) = \frac{\Delta \beta\rho^2 p_0}{(1 - \beta\rho)^2} \tag{24}$$

Corollary (3): Let $\alpha = 0$, $\beta = 1$ and $N \rightarrow \infty$, we obtain a poisson queue with heterogeneity.

Thus from (13), (14), (15) and (16) we have:

$$p_n = \Delta \rho^n p_0, n \geq 1 \tag{25}$$

$$p_0 = \frac{\mu_1 \mu_2 (2 \rho + 1) (1 - \rho)}{\mu_1 \mu_2 (2 \rho + 1) (1 - \rho) + \lambda (\lambda + \mu_1 \pi_2 + \mu_2 \pi_1)}, \tag{26}$$

$$E(n) = \Delta p_0 \left[1 + \frac{\rho}{1 - \rho} + \frac{\rho}{(1 - \rho)^2} \right] \tag{27}$$

$$\text{and } E(m) = \frac{\Delta \rho^2 p_0}{(1 - \rho)^2} \tag{28}$$

This is Krishnamoorthi's work [3].

Corollary (4): Let $\alpha = 0$, $N \rightarrow \infty$ and $\pi_1 = \pi_2$,

we get:

$$p_n = \Delta \rho p_0 (\beta \rho)^{n-2}, n = 2 (1) \dots \quad (29)$$

$$p_0 = \frac{\mu_1 \mu_2 (2 \rho + 1) (1 - \beta \rho)}{\mu_1 \mu_2 (2 \rho + 1) (1 - \beta \rho) + \lambda (\lambda + \mu_2) [1 + (1 - \beta) \rho]} \quad (30)$$

$$E(n) = -\Delta p_0 \left[1 + \frac{\rho}{1 - \beta \rho} + \frac{\rho}{(1 - \beta \rho)^2} \right] \quad (31)$$

$$\text{and } E(m) = \frac{\Delta \beta \rho^2}{(1 - \beta \rho)^2} p_0 \quad (32)$$

Here Δ is different from that of Abou - El - Ata's [1] in Corollary (2) because $e(\pi_1 = \pi_2)$. This is Singh's work [5]. Finally we can deduce Gumble's work [2] if we let $\alpha = 0, \beta = 1, N \rightarrow \infty$ and $\pi_1 = \pi_2 = \frac{1}{2}$. Also Saaty's work [4] if we let $\alpha = 0, \beta = 1, N \rightarrow \infty$ and

$$\pi_1 = \frac{\mu_1}{\mu_1 + \mu_2}, \pi_2 = \frac{\mu_2}{\mu_1 + \mu_2} .$$

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